

Q1				
(i)	$L = \frac{1}{\sigma_1 \sqrt{2\pi}} e^{-\frac{(W_1 - \mu)^2}{2\sigma_1^2}} \cdot \frac{1}{\sigma_2 \sqrt{2\pi}} e^{-\frac{(W_2 - \mu)^2}{2\sigma_2^2}}$ $\ln L = \text{const} - \frac{1}{2\sigma_1^2} (W_1 - \mu)^2 - \frac{1}{2\sigma_2^2} (W_2 - \mu)^2$ $\frac{d \ln L}{d\mu} = \frac{2}{2\sigma_1^2} (W_1 - \mu) + \frac{2}{2\sigma_2^2} (W_2 - \mu)$ $= 0 \Rightarrow \sigma_2^2 W_1 - \sigma_2^2 \mu + \sigma_1^2 W_2 - \sigma_1^2 \mu = 0$ $\Rightarrow \hat{\mu} = \frac{\sigma_2^2 W_1 + \sigma_1^2 W_2}{\sigma_1^2 + \sigma_2^2}$ <p>Check this is a maximum.</p> <p>E.g. $\frac{d^2 \ln L}{d\mu^2} = -\frac{1}{\sigma_1^2} - \frac{1}{\sigma_2^2} < 0$</p>	M1 M1 A1 M1 A1 M1 A1 A1 A1 M1 A1	Product form. Two Normal terms. Fully correct. Differentiate w.r.t. μ . BEWARE PRINTED ANSWER.	11
(ii)	$E(\hat{\mu}) = \frac{\sigma_2^2 \mu + \sigma_1^2 \mu}{\sigma_1^2 + \sigma_2^2} = \mu$ <p>\therefore unbiased.</p>	M1 A1		2
(iii)	$\text{Var}(\hat{\mu}) = \left(\frac{1}{\sigma_1^2 + \sigma_2^2} \right)^2 \cdot (\sigma_2^4 \sigma_1^2 + \sigma_1^4 \sigma_2^2)$ $= \frac{\sigma_1^2 \sigma_2^2}{\sigma_1^2 + \sigma_2^2}$	B1 B1	First factor. Second factor. Simplification not required at this point.	2
(iv)	$T = \frac{1}{2} (W_1 + W_2)$ $\text{Var}(T) = \frac{1}{4} (\sigma_1^2 + \sigma_2^2)$ $\text{Relative efficiency } (y) = \frac{\text{Var}(\hat{\mu})}{\text{Var}(T)}$ $= \frac{\sigma_2^4 \sigma_1^2 + \sigma_1^4 \sigma_2^2}{(\sigma_1^2 + \sigma_2^2)^2} \cdot \frac{4}{\sigma_1^2 + \sigma_2^2}$ $= \frac{4\sigma_1^2 \sigma_2^2}{(\sigma_1^2 + \sigma_2^2)^2}$	B1 M1 M1 A1 A1	Any attempt to compare variances. If correct. BEWARE PRINTED ANSWER.	5
(v)	<p>E.g. consider $\sigma_1^2 + \sigma_2^2 - 2\sigma_1\sigma_2 = (\sigma_1 - \sigma_2)^2 \geq 0$</p> <p>$\therefore$ Denominator \geq numerator, \therefore fraction ≤ 1</p> <p>[Both $\hat{\mu}$ and T are unbiased,] $\hat{\mu}$ has smaller variance than T and is therefore better.</p>	M1 E1 E1 E1		4
				24

Q2	$f(x) = \frac{\lambda^{k+1} x^k e^{-\lambda x}}{k!}, \quad [x > 0 \quad (\lambda > 0, k \text{ integer } \geq 0)]$ <p>Given: $\int_0^\infty u^m e^{-u} du = m!$</p>			
(i)	$M_X(\theta) = E[e^{\theta x}]$ $= \int_0^\infty \frac{\lambda^{k+1}}{k!} x^k e^{-(\lambda-\theta)x} dx$ <p style="text-align: center;">Put $(\lambda - \theta)x = u$</p> $= \frac{\lambda^{k+1}}{k!(\lambda - \theta)^{k+1}} \int_0^\infty u^k e^{-u} du$ $= \left(\frac{\lambda}{\lambda - \theta} \right)^{k+1}$	<p>M1</p> <p>M1</p> <p>M1</p> <p>A1</p> <p>A1</p> <p>A1</p> <p>A1</p>	<p>For obtaining this expression after substitution.</p> <p>Take out constants. (Dep on subst.)</p> <p>Apply "given": integral = k! (Dep on subst.) BEWARE PRINTED ANSWER.</p>	7
(ii)	<p>$Y = X_1 + X_2 + \dots + X_n$</p> <p>By convolution theorem:- mgf of Y is $\{M_X(\theta)\}^n$</p> <p>i.e. $\left(\frac{\lambda}{\lambda - \theta} \right)^{nk+n}$</p> <p>$\mu = M'(0)$</p> $M'(\theta) = \lambda^{nk+n} (-nk - n)(\lambda - \theta)^{-nk-n-1} (-1)$ $\therefore \mu = \frac{nk + n}{\lambda}$ $\sigma^2 = M''(0) - \mu^2$ $M''(\theta) = (nk + n)\lambda^{nk+n} (-nk - n - 1)(\lambda - \theta)^{-nk-n-2} (-1)$ $\therefore M''(0) = (nk + n)(nk + n + 1) / \lambda^2$ $\therefore \sigma^2 = \frac{(nk + n)(nk + n + 1)}{\lambda^2} - \frac{(nk + n)^2}{\lambda^2}$ $= \frac{nk + n}{\lambda^2}$	<p>B1</p> <p>M1</p> <p>A1</p> <p>A1</p> <p>M1</p> <p>A1</p> <p>M1</p> <p>A1</p>		8
(iii)	<p>[Note that $M_Y(t)$ is of the same functional form as $M_X(t)$ with $k + 1$ replaced by $nk + n$, i.e. k replaced by $nk + n - 1$. This must also be true of the pdf.]</p> <p>Pdf of Y is $\frac{\lambda^{nk+n}}{(nk + n - 1)!} \times y^{nk+n-1} \times e^{-\lambda y}$</p> <p>[for $y > 0$]</p>	<p>B1</p> <p>B1</p> <p>B1</p>	<p>One mark for each factor of the expression. Mark for third factor shown here depends on at least one of the other two earned.</p>	3
(iv)	<p>$\lambda = 1, k = 2, n = 5,$ Exact $P(Y > 10) = 0.9165$</p> <p>Use of N(15, 15)</p>	<p>M1</p> <p>M1</p>	<p>Mean. ft (ii).</p> <p>Variance. ft (ii).</p>	

4769

Mark Scheme

June 2006

$P(\text{this} > 10) = P\left(N(0, 1) > \frac{10-15}{\sqrt{15}} = -1.291\right)$ $= 0.9017$	<p>Reasonably good agreement – CLT working for only small n.</p>	<p>A1 A1 E2</p>	<p>c.a.o. c.a.o. (E1, E1) [Or other sensible comments.]</p>	<p>6</p>
				<p>24</p>

Q3				
(i)	$\bar{x} = 36.48 \quad s = 9.6307 \quad s^2 = 92.7507$ $\bar{y} = 45.5 \quad s = 14.8129 \quad s^2 = 219.4218$ <p>Assumptions: Normality of <u>both</u> populations equal variances $H_0 : \mu_A = \mu_B \quad H_1 : \mu_A \neq \mu_B$ Where μ_A, μ_B are the population means.</p> $\text{Pooled } s^2 = \frac{9 \times 92.7507 + 11 \times 219.4218}{20}$ $= \frac{834.756 + 24136.64}{20} = 162.4198$ <p>Test statistic is $\frac{36.48 - 45.5}{\sqrt{162.4198} \sqrt{\frac{1}{10} + \frac{1}{12}}}$</p> $= \frac{-9.02}{5.4568} = -1.653$ <p>Refer to t_{20}. Double tailed 5% point is 2.086. Not significant. No evidence that population mean times differ.</p>	B1 B1 B1 B1 B1 B1 M1 A1 M1 A1 A1 A1	<p>If all correct. [No marks for use of s_n which are 9.1365 and 14.1823 respectively.]</p> <p>Do <u>NOT</u> accept $\bar{X} = \bar{Y}$ or similar.</p> <p>$= (12.7444)^2$</p> <p>No ft from here if wrong. No ft from here if wrong. ft only c's test statistic. ft only c's test statistic.</p>	12
(ii)	<p>Assumption: Normality of underlying population of <u>differences</u>. $H_0 : \mu_D = 0 \quad H_1 : \mu_D > 0$ Where μ_D is the population mean of "before - after" differences.</p> <p>Differences are 6.4, 4.4, 3.9, -1.0, 5.6, 8.8, -1.8, 12.1 $(\bar{x} = 4.8 \quad s = 4.6393)$</p> <p>Test statistic is $\frac{4.8 - 0}{4.6393 / \sqrt{8}}$</p> $= 2.92(64)$ <p>Refer to t_7. Single tailed 5% point is 1.895. Significant. Seems mean is lowered.</p>	B1 B1 B1 M1 M1 A1 M1 A1 A1 A1	<p>Do <u>NOT</u> accept $\bar{D} = 0$ or similar. The "<u>direction</u>" of D must be CLEAR. Allow $\mu_A = \mu_B$ etc.</p> <p>[A1 can be awarded here if NOT awarded in part (i)]. Use of s_n ($=4.3396$) is <u>NOT</u> acceptable, even in a denominator of $\frac{s_n}{\sqrt{n-1}}$</p> <p>No ft from here if wrong. No ft from here if wrong. ft only c's test statistic. ft only c's test statistic.</p>	10
(iii)	The paired comparison in part (ii) eliminates the variability between workers.	E2	(E1, E1)	2
				24

Q4																																																					
(i)	<p>Latin square.</p> <p>Layout such as:</p> <table border="1" style="margin-left: auto; margin-right: auto;"> <thead> <tr> <th colspan="2"></th> <th colspan="5">Locations</th> </tr> <tr> <th colspan="2"></th> <th>1</th> <th>2</th> <th>3</th> <th>4</th> <th>5</th> </tr> </thead> <tbody> <tr> <td></td> <td>I</td> <td>A</td> <td>B</td> <td>C</td> <td>D</td> <td>E</td> </tr> <tr> <td>Surf</td> <td>II</td> <td>B</td> <td>C</td> <td>D</td> <td>E</td> <td>A</td> </tr> <tr> <td>-aces</td> <td>III</td> <td>C</td> <td>D</td> <td>E</td> <td>A</td> <td>B</td> </tr> <tr> <td></td> <td>IV</td> <td>D</td> <td>E</td> <td>A</td> <td>B</td> <td>C</td> </tr> <tr> <td></td> <td>V</td> <td>E</td> <td>A</td> <td>B</td> <td>C</td> <td>D</td> </tr> </tbody> </table>			Locations							1	2	3	4	5		I	A	B	C	D	E	Surf	II	B	C	D	E	A	-aces	III	C	D	E	A	B		IV	D	E	A	B	C		V	E	A	B	C	D	B1		
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(ii)	<p>$X_{ij} = \mu + \alpha_i + e_{ij}$</p> <p>$\mu$ = population grand mean for whole experiment.</p> <p>α_i = population mean amount by which the i^{th} treatment differs from μ.</p> <p>e_{ij} are experimental errors ~ ind $N(0, \sigma^2)$.</p>	B1																																																			
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(iii)	<p>Totals are: 322, 351, 307, 355, 291 (each from sample of size 5) Grand total: 1626</p> <p>"Correction factor" $CF = \frac{1626^2}{25} = 105755.04$</p> <p>Total SS = 106838 – CF = 1082.96</p> <p>Between paints SS = $\frac{322^2}{5} + \dots + \frac{291^2}{5} - CF$ = 106368 – CF = 612.96</p> <p>Residual SS (by subtraction) = 1082.96 – 612.96 = 470.00</p> <table border="1" style="margin-left: auto; margin-right: auto;"> <thead> <tr> <th>Source of variation</th> <th>SS</th> <th>df</th> <th>MS</th> </tr> </thead> <tbody> <tr> <td>Between paints</td> <td>612.96</td> <td>4</td> <td>153.24</td> </tr> <tr> <td>Residual</td> <td>470.00</td> <td>20</td> <td>23.5</td> </tr> <tr> <td>Total</td> <td>1082.96</td> <td>24</td> <td></td> </tr> </tbody> </table> <p>MS ratio = $\frac{153.24}{23.5} = 6.52$</p>	Source of variation	SS	df	MS	Between paints	612.96	4	153.24	Residual	470.00	20	23.5	Total	1082.96	24																																					
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4769

Mark Scheme

June 2006

	<p>Refer to $F_{4, 20}$</p> <p>Upper 5% point is 2.87 Significant.</p> <p>Seems performances of paints are not all the same.</p>	<p>M1</p> <p>A1</p> <p>A1</p> <p>A1</p>	<p>No ft if wrong. But allow ft of wrong d.o.f. above.</p> <p>No ft if wrong.</p> <p>ft only c's test statistic and d.o.f.'s.</p> <p>ft only c's test statistic and d.o.f.'s.</p>	<p>12</p>
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